ISSN 0256 - 422 X

EXPLICIT EXPRESSIONS FOR MOMENTS OF GENERALIZED GAMMA ORDER STATISTICS

SARALEES NADARAJAH

School of Mathematics, University of Manchester, Manchester M13 9PL, UK Email: saralees.nadarajah@manchester.ac.uk

SUMMARY

Explicit closed form expressions for moments of order statistics from the generalized gamma distribution are derived using Lauricella function of type A.

Keywords and phrases: Gamma distribution; Generalized gamma distribution; Lauricella function of type A; Order statistics.

AMS Classification: 33C90; 62E99.

1 Introduction

Suppose X_1, X_2, \ldots, X_n is a random sample from the generalized gamma distribution given by the probability density function (pdf):

$$f(x) = \frac{cx^{c\alpha-1}\exp\left(-x^c\right)}{\Gamma\left(\alpha\right)}$$
(1.1)

for x > 0, $\alpha > 0$ and c > 0. Let $X_{1:n} < X_{2:n} < \cdots < X_{n:n}$ denote the corresponding order statistics. There has been some work relating to moments of the generalized gamma order statistics $X_{r:n}$, see e.g. Khan and Khan (1983) and Thomas (1996). However, all of the work that we are aware of express $E(X_{r:n}^k)$ in terms of recurrence relations, numerical tables or multiple infinite series of the form

$$E\left(X_{r:n}^{k}\right) = \sum_{m_{1}=0}^{\infty} \cdots \sum_{m_{j}=0}^{\infty} A\left(m_{1}, \dots, m_{j}, r, n, k\right)$$

$$(1.2)$$

for some function $A(\cdot)$ and j = j(r, n, k). These representations have little practical appeal because, for example, the computation of the *j*-fold infinite sum in (1.2) will become prohibitive as *j* gets large. Even for moderate *j*, one would have to code in (1.2) and this will be a waste of man power as well as computer time.

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In this note, for the first time, we derive expressions for $E(X_{r:n}^k)$ that are finite sums of a well known special function – namely, the Lauricella function of type A (Exton, 1978) defined by

$$F_{A}^{(n)}(a, b_{1}, \dots, b_{n}; c_{1}, \dots, c_{n}; x_{1}, \dots, x_{n}) = \sum_{m_{1}=0}^{\infty} \cdots \sum_{m_{n}=0}^{\infty} \frac{a_{m_{1}+\dots+m_{n}}(b_{1})_{m_{1}}\cdots(b_{n})_{m_{n}}}{(c_{1})_{m_{1}}\cdots(c_{n})_{m_{n}}} \frac{x_{1}^{m_{1}}\cdots x_{n}^{m_{n}}}{m_{1}!\cdots m_{n}!},$$
(1.3)

where $(f)_k = f(f+1)\cdots(f+k-1)$ denotes the ascending factorial. Numerical routines for the direct computation of (1.3) are widely available, see e.g. Mathematica and Exton (1978).

2 Explicit Expression for $E(X_{r:n}^k)$

If X_1, X_2, \ldots, X_n is a random sample from (1.1) then it is well known that the pdf of $Y = X_{r:n}$ is given by

$$f_Y(y) = \frac{n!}{(r-1)!(n-r)!} \left\{ F(y) \right\}^{r-1} \left\{ 1 - F(y) \right\}^{n-r} f(y)$$

for r = 1, 2, ..., n, where $F(\cdot)$ is the cumulative distribution function (cdf) corresponding to (1.1) given by

$$F(y) = \frac{\gamma(\alpha, y^c)}{\Gamma(\alpha)},$$

where $\gamma(\cdot, \cdot)$ denotes the incomplete gamma function defined by

$$\gamma(\alpha, x) = \int_0^x t^{\alpha - 1} \exp\left(-t\right) dt$$

Thus, the kth moment of $X_{r:n}$ can be expressed as

$$E\left(X_{r:n}^{k}\right) = \frac{cn!}{(r-1)!(n-r)!\left\{\Gamma(\alpha)\right\}^{n}} \\ \times \int_{0}^{\infty} y^{k+c\alpha-1} \exp\left(-y^{c}\right) \left\{\gamma\left(\alpha,y^{c}\right)\right\}^{r-1} \left\{\Gamma(\alpha) - \gamma\left(\alpha,y^{c}\right)\right\}^{n-r} dy \\ = \frac{cn!}{(r-1)!(n-r)!\left\{\Gamma(\alpha)\right\}^{n}} \int_{0}^{\infty} y^{k+c\alpha-1} \exp\left(-y^{c}\right) \\ \times \sum_{l=0}^{n-r} {n-r \choose l} \left\{\Gamma(\alpha)\right\}^{n-r-l} (-1)^{l} \left\{\gamma\left(\alpha,y^{c}\right)\right\}^{r+l-1} dy \\ = \frac{cn!}{(r-1)!(n-r)!} \sum_{l=0}^{n-r} (-1)^{l} {n-r \choose l} \left\{\Gamma(\alpha)\right\}^{-r-l} \\ \times \int_{0}^{\infty} y^{k+c\alpha-1} \exp\left(-y^{c}\right) \left\{\gamma\left(\alpha,y^{c}\right)\right\}^{r+l-1} dy \\ = \frac{cn!}{(r-1)!(n-r)!} \sum_{l=0}^{n-r} (-1)^{l} {n-r \choose l} \left\{\Gamma(\alpha)\right\}^{-r-l} dy$$

$$(2.1)$$

Using the series expansion

$$\gamma(\alpha, x) = x^{\alpha} \sum_{m=0}^{\infty} \frac{(-x)^m}{(\alpha+m)m!},$$

the integral I(l) in (2.1) can be expressed as

$$I(l) = \int_{0}^{\infty} y^{k+c\alpha-1} \exp\left(-y^{c}\right) \left\{ y^{c\alpha} \sum_{m=0}^{\infty} \frac{(-y^{c})^{m}}{(\alpha+m)m!} \right\}^{r+l-1} dy$$

$$= \int_{0}^{\infty} \sum_{m_{1}=0}^{\infty} \cdots \sum_{m_{r+l-1}=0}^{\infty} \frac{(-1)^{m_{1}+\dots+m_{r+l-1}}y^{k+c\alpha(r+l)+c(m_{1}+\dots+m_{r+l-1})-1} \exp\left(-y^{c}\right)}{(\alpha+m_{1})\cdots(\alpha+m_{r+l-1})m_{1}!\cdots m_{r+l-1}!} dy$$

$$= \sum_{m_{1}=0}^{\infty} \cdots \sum_{m_{r+l-1}=0}^{\infty} \frac{(-1)^{m_{1}+\dots+m_{r+l-1}}}{(\alpha+m_{1})\cdots(\alpha+m_{r+l-1})m_{1}!\cdots m_{r+l-1}!} \exp\left(-y^{c}\right) dy$$

$$= \sum_{m_{1}=0}^{\infty} \cdots \sum_{m_{r+l-1}=0}^{\infty} \frac{(-1)^{m_{1}+\dots+m_{r+l-1}}\Gamma\left(k/c+\alpha(r+l)+m_{1}+\dots+m_{r+l-1}\right)}{c(\alpha+m_{1})\cdots(\alpha+m_{r+l-1})m_{1}!\cdots m_{r+l-1}!}. \quad (2.2)$$

Using the fact $(f)_k = \Gamma(f+k)/\Gamma(f)$ and the definition in (1.3), one can reexpress (2.2) as

$$I(l) = c^{-1} \alpha^{1-r-l} \Gamma \left(k/c + \alpha(r+l) \right) \\ \times F_A^{(r+l-1)} \left(k/c + \alpha(r+l), \alpha, \dots, \alpha; \alpha+1, \dots, \alpha+1; -1, \dots, -1 \right).$$
(2.3)

Combining (2.1) and (2.3), we obtain the expression

$$E\left(X_{r:n}^{k}\right) = \frac{n!}{(r-1)!(n-r)!} \sum_{l=0}^{n-r} (-1)^{l} {\binom{n-r}{l}} \left\{\Gamma(\alpha)\right\}^{-r-l} \alpha^{1-r-l} \Gamma\left(k/c + \alpha(r+l)\right) \\ \times F_{A}^{(r+l-1)} \left(k/c + \alpha(r+l), \alpha, \dots, \alpha; \alpha+1, \dots, \alpha+1; -1, \dots, -1\right) (2.4)$$

Note that (2.4) is a finite sum of the Lauricella function of type A.

3 Conclusions

We have derived an expression for moments of the generalized gamma order statistics as a finite sum of a well known special function. This expression is simpler than previously known work.

References

- Exton, H. (1978). Handbook of Hypergeometric Integrals: Theory, Applications, Tables, Computer Programs. Halsted Press, New York.
- [2] Khan, A. H. and Khan, R. U. (1983). Recurrence relations between moments of order statistics from generalized gamma distribution. *Journal of Statistical Research*, 17, 75– 82.
- [3] Thomas, P. Y. (1996). On the moments of extremes from generalized gamma distribution. Communications in Statistics—Theory and Methods, 25, 1825–1836.